

# Dependable and Reliable Cloud-Based Systems Using Multiple Software Spare Components

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**Abstract**—Cloud computing relies on a set of service components running on a service provider datacenter to achieve specific tasks. A trusted cloud-based software system is a highly dependable, reliable, available and predictable advanced computing system with guaranteed Quality of Service (QoS). Due to well-established studies and practices on hardware reliability, software faults have become the major factor of system failures in cloud-based systems. In this paper, we introduce a scheme of developing dependable and reliable cloud-based systems using multiple software spare components. We address the software-aging phenomenon in cloud computing, where the reliability of a software component decreases along the time. To counteract the software aging issue, we propose a mechanism to maintain the system reliability above a predefined safety threshold using software rejuvenation schedules. The calculation of system reliability is based on an extended Dynamic Fault Tree (DFT) model of cloud-based systems with Software SPare (SSP) gates. We verify our approach using Continuous Time Markov Chain (CTMC) for the case of constant failure rates, and provide a case study of a cloud-based system to show the detailed procedure as well as the feasibility of our approach.

**Keywords**—Software aging; hot software spare; cold software spare; software spare gate; reliability analysis; dynamic fault tree; software rejuvenation schedule

## I. INTRODUCTION

As the cloud computing paradigm continues to grow along with the rapid computing technology advancement, cloud-based services are increasingly being used in many different areas such as healthcare, public transportation, mobile cloud computing and many more. A trusted cloud-based system is a highly dependable, reliable, available and predictable advanced computing system with guaranteed Quality of Service (QoS). The QoS of a computer-based system has been widely researched to maintain fault-tolerant hardware, secure, available and reliable software resources for client consumption. However, system outage of cloud-based systems is still common despite the well-established fault-tolerant techniques for hardware [1]. Software related faults of cloud-based systems due to the software-aging phenomenon [2] have become one of the major obstacles to achieving high fault tolerance and system reliability. Therefore, we were motivated to resolve the software aging related issues in cloud computing in order to maintain high dependability and reliability of

cloud-based systems. In this work, we perform system reliability analysis from the perspective of Software Reliability Engineering (SRE). Early SRE focused on the analysis of software defects and bugs including Bohrbugs and Heisenbugs; while recently, the concept of software aging was introduced [3], taking into account the growing usage of cloud computing and the increasing workload that impacts the reliability of cloud-based systems. The software aging phenomenon is due to the degradation of system resources used by a software system until failure, which is caused by many factors such as memory bloating, memory leaks, data corruption, unreleased file-locks, unterminated threads, accumulation of round-off errors, and storage and space fragmentation [4]. To counteract the software aging problem, software rejuvenation has been proposed as a solution for achieving high fault tolerance in software-based systems [5]. Software rejuvenation can be done in many different ways, where the simplest one is to restart the application that causes the aging problem, or to reboot the whole system.

Correctly measuring the reliability of a cloud-based system is critical to avoid software failures due to the software-aging phenomenon. In this paper, we extend our former analytical-based approach to deriving the reliability function of a hot spare gate with a single hot standby spare [6]. In our new approach, multiple software spares are used for critical software components in cloud computing. The significance of our new approach is described as follows. First, as our former approach does not scale well for multiple hot spare parts, when a second hot spare is added into the system design, the analytical approach becomes non-trivial for the formalization of the analysis process. Second, it is useful and important to understand how a rejuvenation schedule may be affected by multiple hot software spares. We show the technical details for deriving the reliability function for a Software SPare (SSP) gate with two hot spare components, and use Continuous Time Markov Chain (CTMC) to verify the correctness of our approach for constant failure rates. A practical case study of a cloud-based system with two hot spares for two critical software components has been provided. In the case study, we assume a reliability threshold for triggering the software rejuvenation process, and based on the cloud-based system reliability analysis, we derive a software rejuvenation schedule to improve system reliability, dependability and availability.

## II. RELATED WORK

Considerable research has been conducted on software aging and software rejuvenation to achieve high fault tolerance in software systems. There are mainly two categories of approaches to predicting software rejuvenation schedules, namely measurement-based and analytical-based approaches [7]. Measurement-based approach uses statistical analysis for the measured data of resource degradation that leads to software aging faults. A monitoring program collects the data, and analyzes them in order to estimate the degradation level. The rejuvenation process is triggered based on a predefined degradation threshold. Grottko *et al.* analyzed the resource degradation in a web server subject to injected workload [8]. The existence of monotonic trends was tested in time series, where these trends are indications of the software aging issues. Machida *et al.* detected software aging by applying Mann-Kendall test that is based on traces of computer system metrics [9]. Guo *et al.* established a trend prediction method to uncover software aging based on the quality of user requests [10]. Measurement-based approaches are feasible ways of predicting software aging, but they are quite inaccurate, and expensive in computational requirements due to the processing of large amounts of system data. Therefore, they are inefficient approaches in practical usage. However, when we use the time-to-failure distribution for data fitting and the calculation of system reliability, the estimated distribution from measurements can be useful in our proposed analytical-based approach [11].

On the other hand, in analytical-based approaches, we first have to assume failure time distributions for the components or the systems subject to software aging, and then schedule software rejuvenation processes at fixed interval based on the analytical results of the system reliability and availability. Bobbio *et al.* suggested a fine-grained software degradation model for optimal rejuvenation scheduling [12], by identifying the system current degradation level that outlines two different strategies of rejuvenation policies. Vaidyanathan *et al.* worked on an analytical model for software systems that uses inspection-based software rejuvenation [13]. They showed the advantages of inspection-based maintenance over non-inspection-based maintenance using Semi-Markov modeling. Koutras and Platis addressed a software rejuvenation technique for cluster systems, where rejuvenation can be carried out when node-deployed software starts to experience degradation, and thus an unscheduled reboot may be avoided [14]. Despite the fact that the above approaches introduced different models for software rejuvenation, they cannot be used to model dynamic behaviors such as sparing and dynamic relationships. Unlike the existing analytical-based approaches, our method studies the dynamic behaviors of software components in cloud-based software systems, namely, the standby software sparing components, and provides a novel analytical approach for reliability analysis.

Moreover, in the context of standby systems, there are four categories of evaluation methods for analyzing standby systems, namely simulations, state-space based methods, analytical/combinatorial approach, and numerical approach. As mentioned previously, simulation methods are expensive in terms of computations, and can only lead to approximate

results [15]. Markov-based methods are state-space oriented [16], while non-Markovian models [17] are powerful in dynamic modeling. However, Markov-based models are limited to exponential failure distributions, and both of the approaches experience the state-space explosion problem when modeling large systems. Analytical approaches, such as minimal cut sets or sequences [18], and sequential decision diagrams [19] are limited to modeling complex behaviors with various time-to-failure distribution types. In our approach, we propose an extended DFT to model the reliability of cloud-based systems. We introduce an analytical-based approach to analyzing the extended DFT model for reliability calculation. Our approach does not suffer from the state-space explosion problem as it is compositional, where a DFT is decomposed into subtrees, and the system reliability is calculated by joining the reliabilities of the subtrees. Finally, numerical methods have been used as an iterative way for analyzing various designs of standby systems with a discrete approximation of time-to-failure distributions [20]. It is potential that a numerical method as demonstrated in previous work could be useful in our proposed analytical-based approach for estimating the time-to-failure distribution.

Finally, there is also some previous work on virtualized datacenter and cloud-based systems. Machida *et al.* proposed an availability model for virtualized systems with time-based rejuvenation using Petri-nets and a gradient search method [21]. Thein *et al.* modeled the availability of application servers, and they showed the high-availability cluster failover, combining virtualization and software rejuvenation [22]. However, none of the above approaches addressed explicitly the reliability analysis for software rejuvenation scheduling. In our approach, we analyze system reliability using an extended DFT model and use the proposed analytical approach to estimate rejuvenation schedules that satisfy predefined reliability requirements for cloud-based systems.

## III. REJUVENATION IN THE CLOUD USING SOFTWARE SPARES

Virtualization allows multiple clients to share a physical machine's resources using virtual machines (VM). To maintain high fault tolerance of a cloud-based system subject to software aging, we employ software rejuvenation and standby sparing for software redundancy to ensure service continuity. Different from physical machines, VMs are stored as images, which can be easily created, managed, and destroyed, making them very suitable for disaster recovery and disaster prevention. Comparing to hardware spares, using sparing VMs for disaster prevention and software rejuvenation could be a very inexpensive and effective way to restore the high performance of a cloud-based software system.

To achieve a reliable and zero-downtime rejuvenation, we define two types of VM spares, namely Hot Software Spare (HSS) and Cold Software Spare (CSS). In the context of cloud-based systems, an HSS is a hot standby VM instance that can be instantly available when a primary component fails. Despite the fact that an HSS is running alongside a primary component, it is not sharing any workload or processing any requests. Therefore, an HSS is operated using much less CPU power, but can be scaled automatically to meet the workload requirements when a primary component fails.

Critical data in an HSS is mirrored in near real time from the primary VM instance, e.g., in the range of 200  $\mu$ s, to ensure high fault tolerance. The failure rate of an HSS is much less than that of a primary component as an HSS is not subject to aging-related bugs. This makes a software-defined HSS differ significantly from a hardware-based Hot SPare (HSP) because, with physical wearout, an HSP may have the same failure rate as a primary hardware component. On the other hand, a CSS refers to a software component that is available as an image of a VM, and can be replicated and deployed as a primary component or a HSS component. As an inactive VM instance, a CSS is mirrored for its critical data based on a specified schedule with most of the time being cold standby. Therefore, the reliability of a CSS is nearly perfect, which can be reasonably assumed never to fail. The recovery time using a CSS is usually in the range of minutes up to two hours; while the cost of a CSS is its storage and very little CPU resource consumption. A CSS can be rapidly deployed, which makes it quite different from a hardware-based Cold SPare (CSP) that is much expensive and requires manual configuration when a primary one fails.

Software rejuvenation techniques have been used to prevent the occurrence of aging-related software failures by proactively resetting a system's internal state to its initial condition. In this work, we adopt an easy way of software rejuvenation by rebooting the system according to a defined schedule. In cloud computing, we can start a new VM to replace an old one that has demonstrated unsatisfactory system performance. To render the fault tolerant of the critical components and minimize the frequency of the rejuvenation events, each critical primary component is equipped by at least two HSSs and one CSS. The only CSS is needed for the rejuvenation process – it can be replicated for all currently deployed software components including the primary one and the HSSs. A newly deployed component must wait until the old ones have finished processing their remaining requests before they can be destroyed.

In our approach, a rejuvenation process is triggered when the reliability of a system component or the whole system reaches a predefined threshold. Similar to [6], we assume the rejuvenation process (Phase 1) takes about 30 minutes, with sufficient time to start a CSS and complete all remaining requests before Phase 2 starts. As a CSS never fails, we only consider the primary component and its HSSs when calculating the system reliability. In addition, two scenarios are investigated for the rejuvenation procedure. One scenario, called *system-specific* rejuvenation, is to rejuvenate the whole system when the system reliability reaches a threshold. The second scenario is a *component-specific* one, in which the critical component with the lowest reliability is rejuvenated when the system reliability reaches a threshold. As shown in a case study, the component-specific rejuvenation demonstrates certain advantages over the system-specific approach.

#### IV. RELIABILITY MODELING AND ANALYSIS

Dynamic Fault Tree (DFT) extends the concept of static fault tree and introduces new modeling capabilities for spare components, functional dependency, and failure sequence dependency. In this paper, we further extend DFT for

modeling software spare components in cloud-based systems with software aging phenomenon.

##### A. SSP Gate for Cloud-Based Systems with Two Hot Spares

Figure 1 shows a SSP gate with one primary component  $P$  and two HSS components  $H_1$  and  $H_2$ . The primary component is initially powered on, but when it fails, it is replaced by an alternate spare following an enumeration sequence. Therefore, a SSP gate fails only when the primary component and all the alternate HSS components fail. Suppose the constant failure rates of components  $P$ ,  $H_1$ , and  $H_2$  are  $\lambda_P$ ,  $\lambda_{H1}$ , and  $\lambda_{H2}$ , respectively. When  $P$  fails,  $H_1$  takes the lead to replace  $P$  as  $H_1^*$ , with  $\lambda_{H1^*} \geq \lambda_{H1}$  due to the software aging phenomenon, when it takes the full workload. The same thing happens to  $H_2$  when  $H_1^*$  fails –  $H_2$  replaces  $H_1^*$  as  $H_2^*$  with  $\lambda_{H2^*} \geq \lambda_{H2}$ . Note that  $\lambda_{H^*}$  and  $\lambda_P$  do not have to be equal because  $P$  and  $H$  may have different configurations. In addition, we designate  $\tau_1$ ,  $\tau_2$ , and  $\tau_3$  as the time to failure of  $P$ ,  $H_1$  and  $H_2$ , respectively.

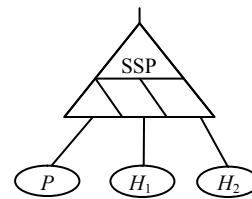


Fig. 1. An SSP gate with a primary component and two HSSs

To derive the reliability function of a SSP gate with two hot spares, we identify all the possible events when a SSP gate fails according to component failure sequence. We denote the event “component  $X$  fails before component  $Y$ ” as  $X \prec Y$ , and summarize six disjoint events  $e_i$  where  $1 \leq i \leq 6$ , as in Fig. 2.

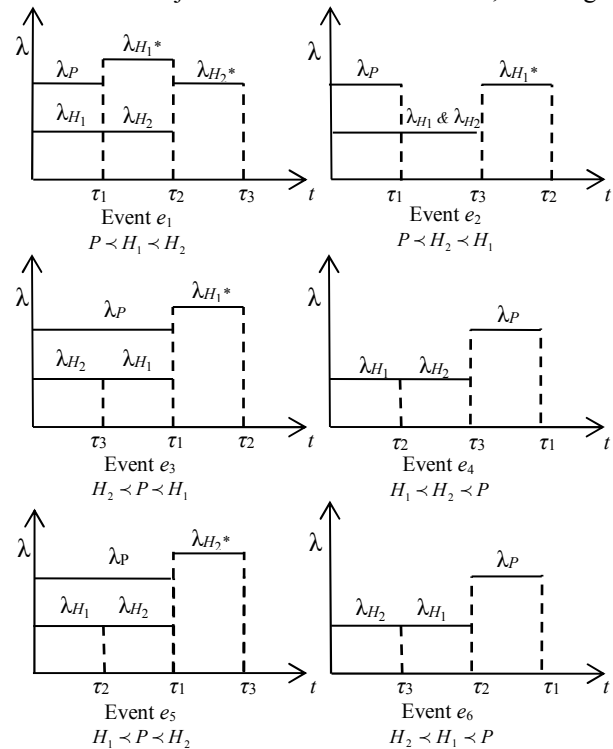


Fig. 2. Six events for the failure of an SSP gate with two HSSs

Let event  $A$  be the failure of an SSP gate at time  $t$ . We can calculate the probability of event  $A$  as in Eq. (1):

$$\Pr(A) = \sum_{i=1}^6 \Pr(A|e_i) \Pr(e_i) = \sum_{i=1}^6 \Pr(A \cap e_i) \quad (1)$$

It is worth noting that when event  $e_i$  happens, the SSP gate also fails. Therefore, event  $A$  always happens with some event  $e_i$ . Thus, Eq. (1) can be simplified as in Eq. (2).

$$\Pr(A) = \sum_{i=1}^6 \Pr(e_i) \quad (2)$$

**Event  $e_1$ :**  $P$  fails before  $H_1$ , and  $H_1$  fails before  $H_2$ , denoted as  $P \prec H_1 \prec H_2$ . In this case, it is guaranteed that  $H_1$  does not fail during  $(0, \tau_1]$ , and  $H_2$  does not fail during  $(0, \tau_2]$ . After  $P$  fails,  $H_1$  takes over the workload and becomes  $H_1^*$ , also after  $H_1^*$  fails,  $H_2$  takes over the workload and becomes  $H_2^*$ . Intuitively, the unreliability function  $U(t)$  of the SSP gate, i.e., the probability that the SSP gate fails during  $(0, t]$ , can be calculated as in Eq. (3).

$$\Pr_{P \prec H_1 \prec H_2} (T \leq t) = \int_0^t \int_0^{\tau_1} \int_0^{\tau_2} (\lambda_P e^{-\lambda_P \tau_1}) (\lambda_{H_1^*} e^{-\lambda_{H_1^*} \tau_2}) (\lambda_{H_2^*} e^{-\lambda_{H_2^*} \tau_3}) d\tau_3 d\tau_2 d\tau_1 \quad (3)$$

However, Eq. (3) only works when  $\lambda_{H_1^*} = \lambda_{H_1}$ . As shown in previous work [6], when  $\lambda_{H_1^*} > \lambda_{H_1}$ , the integration of the probability density function (*pdf*) of  $H_1^*$  from  $\tau_1$  to  $t$  does not give the correct unreliability of the component at time  $t$ , as it incorrectly assumes that component  $H_1$  behaves as  $H_1^*$  starting from time 0. Since the component actually behaves as  $H_1$  during  $(0, \tau_1]$ , the unreliability of  $H_1^*$  at time  $\tau_1$  equals the unreliability of  $H_1$  at  $\tau_1$  rather than the unreliability calculated by the integration of the *pdf* of  $H_1^*$  from 0 to  $\tau_1$ . This is to ensure the unreliability continuity for  $H_1$  before and after it serves as a primary component  $H_1^*$ . By calculating a new starting integration time  $\tau_{H_1^*}$  for  $H_1^*$ , we take into consideration that  $\tau_2$ , originally the failure of component  $H_1$ , is shifted to the left by  $(\tau_1 - \tau_{H_1^*})$ . As a result, when we consider the failure of  $H_1^*$ , we must add  $(\tau_1 - \tau_{H_1^*})$  to  $\tau_2$  since  $H_2^*$  is activated based on the original non-shifted failure time variable  $\tau_2$  of  $H_1$ . Therefore, the value of  $\tau_2$  after the adjustment is given as  $\tau_{2\text{actual}} = \tau_{2\text{shifted}} + (\tau_1 - \tau_{H_1^*}) = \tau_2 + (\tau_1 - \tau_{H_1^*})$ , as shown in Fig. 3. As a rule of thumb, in the case of  $P \prec H_1 \prec H_2 \dots \prec H_i$ , where  $i > 1$  ( $\tau_1$  does not get shifted since it is the failure time of  $P$ , and  $P$  always acts as a primary component), when a component  $H_i^*$  acts as a primary one, its actual time to failure equals  $\tau_{(i+1)} + (\tau_i - \tau_{H_i^*})$ . This observation and adjustment is critical for yielding the correct reliability function.

Hot spare  $H_1$  or the first HSS has been studied in previous work [6] yielding  $\tau_{H_1^*} = (\lambda_{H_1} / \lambda_{H_1^*}) \tau_1$ . In regards to the second HSS  $H_2$ , it is guaranteed that  $H_2$  does not fail during  $(0, \tau_2]$ . After  $H_1^*$  fails,  $H_2$  takes over the workload and becomes  $H_2^*$ . Since the component actually behaves as  $H_2$  during  $(0, \tau_2]$ , the unreliability of  $H_2^*$  at time  $\tau_2$  equals the unreliability of  $H_2$  at  $\tau_2$  rather than the unreliability calculated by the integration of the *pdf* of  $H_2^*$  from 0 to  $\tau_2$ . This requires us to calculate a new starting integration time  $\tau_{H_2^*}$  for  $H_2^*$  such that the unreliability of  $H_2^*$  at  $\tau_{H_2^*}$  is equal to the unreliability of  $H_2$  at  $\tau_2$ . As the *pdfs* of  $H_2$  and  $H_2^*$  are  $f(\tau) = \lambda_{H_2} e^{-\lambda_{H_2} \tau}$  and  $f(\tau) = \lambda_{H_2^*} e^{-\lambda_{H_2^*} \tau}$ , respectively, such a relationship between  $H_2$  and  $H_2^*$  can be described as in Eq. (4), taking into account the adjustment of  $\tau_2$ , i.e., the time to failure of  $H_1^*$ .

$$\int_0^{\tau_{H_2^*}} \lambda_{H_2^*} e^{-\lambda_{H_2^*} \tau_2} d\tau_2 = \int_0^{\tau_2 + (\tau_1 - \tau_{H_1^*})} \lambda_{H_2} e^{-\lambda_{H_2} \tau_2} d\tau_2 \quad (4)$$

Solving Eq. (4), we have  $\tau_{H_2^*} = \frac{\lambda_{H_2}}{\lambda_{H_2^*}} (\tau_2 + (\tau_1 - \tau_{H_1^*}))$ . Since  $H_2^*$  fails during a period of time  $(t - \tau_2)$ , the integration range for  $H_2^*$  now becomes  $[\tau_{H_2^*}, t - (\tau_2 + (\tau_1 - \tau_{H_1^*})) + \tau_{H_2^*}]$ , as illustrated in Fig. 3. The probability of the event  $P \prec H_1 \prec H_2$ , i.e.,  $\Pr(e_1)$ , can be calculated as in Eq. (5).

$$\Pr_{P \prec H_1 \prec H_2} (T \leq t) = \int_0^t \int_{\tau_{H_1^*}}^{t - (\tau_1 + \tau_{H_1^*}) - (t - (\tau_2 + (\tau_1 - \tau_{H_1^*})) + \tau_{H_2^*})} (\lambda_P e^{-\lambda_P \tau_1}) (\lambda_{H_1^*} e^{-\lambda_{H_1^*} \tau_2}) (\lambda_{H_2^*} e^{-\lambda_{H_2^*} \tau_3}) d\tau_3 d\tau_2 d\tau_1 \quad (5)$$

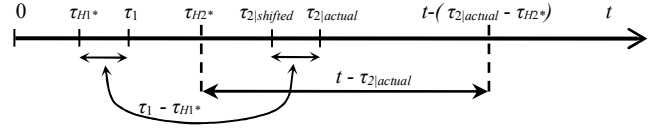


Fig. 3. Failure time of  $H_1$  and  $H_2$  for reliability analysis in event  $e_1$

Eq. (5) can be simplified with two substitutions  $w(\tau_2) = \tau_2 + \tau_1 - \tau_{H_1^*}$  and  $z(\tau_3) = \tau_3 + w - \tau_{H_2^*}$ , which results in Eq. (6), where  $z$  and  $w$  can be replaced by  $\tau_2$  and  $\tau_3$ , respectively.

$$\Pr_{P \prec H_1 \prec H_2} (T \leq t) = \int_0^t \int_{\tau_{H_1^*}}^t (\lambda_P e^{-\lambda_P \tau_1}) (\lambda_{H_1^*} e^{-\lambda_{H_1^*} [w - \tau_1 + \tau_{H_1^*}]} (\lambda_{H_2^*} e^{-\lambda_{H_2^*} [z - w - \frac{\lambda_{H_2}}{\lambda_{H_2^*}}]}) dz dw d\tau_1 \quad (6)$$

**Event  $e_2$ :**  $P \prec H_2 \prec H_1$ , this is where  $P$  fails first then  $H_2$  fails as a spare before  $H_1^*$  fails. The failure of  $H_2$  is independent of  $H_1^*$ , and the failure of  $H_1^*$  depends on  $P$  failure but not on  $H_2$ 's failure. The integration of  $H_1^*$  requires computing  $\tau_{H_1^*}$ , which is based on  $\tau_1$  by moving the integration limit from  $\tau_{H_1^*}$  to  $\tau_{H_1^*} + (\tau_3 - \tau_1)$ , resulting in Eq. (7).

$$\Pr_{P \prec H_2 \prec H_1} (T \leq t) = \int_0^t \int_{\frac{\lambda_{H_1^*} \tau_1 + (\tau_3 - \tau_1)}{\lambda_{H_1^*}}}^{t - (\tau_1 - \tau_{H_1^*})} (\lambda_P e^{-\lambda_P \tau_1}) (\lambda_{H_1^*} e^{-\lambda_{H_1^*} \tau_2}) (\lambda_{H_2} e^{-\lambda_{H_2} \tau_3}) d\tau_2 d\tau_3 d\tau_1 \quad (7)$$

**Event  $e_3$ :**  $H_2 \prec P \prec H_1$ , this is where  $H_2$  fails first as a spare, then  $P$  fails, and finally  $H_1$  fails as  $H_1^*$ . Note that the complexity is similar to one spare SSP gate  $P \prec H$ . The probability that the SSP gate fails is calculated as in Eq. (8).

$$\Pr_{H_2 \prec P \prec H_1} (T \leq t) = \int_0^t \int_{\frac{\lambda_{H_1^*} \tau_1}{\lambda_{H_1^*}}}^{t - (\tau_1 - \tau_{H_1^*})} (\lambda_P e^{-\lambda_P \tau_1}) (\lambda_{H_1^*} e^{-\lambda_{H_1^*} \tau_2}) (\lambda_{H_2} e^{-\lambda_{H_2} \tau_3}) d\tau_2 d\tau_3 d\tau_1 \quad (8)$$

**Event  $e_4$ :**  $H_1 \prec H_2 \prec P$ ,  $H_1$  and  $H_2$  fail as spares before  $P$  fails, similar to one spare SSP gate, where it is guaranteed that  $P$  does not fail during  $(0, \tau_3]$ . The probability that the SSP gate fails during  $(0, t]$  can be calculated as in Eq. (9).

$$\Pr_{H_1 \prec H_2 \prec P} (T \leq t) = \int_0^t \int_{\tau_3}^t (\lambda_P e^{-\lambda_P \tau_1}) (\lambda_{H_1} e^{-\lambda_{H_1} \tau_2}) (\lambda_{H_2} e^{-\lambda_{H_2} \tau_3}) d\tau_2 d\tau_3 d\tau_1 \quad (9)$$

**Event  $e_5$ :**  $H_1 \prec P \prec H_2$ , similar to event  $e_3$ , this is where  $H_1$  fails first as a spare, then  $P$  and  $H_2$  as  $H_2^*$ . Note that the complexity is similar to one spare SSP gate  $P \prec H$ . The probability that the SSP gate fails is calculated as in Eq. (10).

$$\Pr_{H_1 \prec P \prec H_2} (T \leq t) = \int_0^t \int_{\frac{\lambda_{H_2} \tau_2}{\lambda_{H_2^*}}}^{t - (\tau_1 - \tau_{H_1^*})} (\lambda_P e^{-\lambda_P \tau_1}) (\lambda_{H_1} e^{-\lambda_{H_1} \tau_2}) (\lambda_{H_2^*} e^{-\lambda_{H_2^*} \tau_3}) d\tau_2 d\tau_3 d\tau_1 \quad (10)$$

**Event  $e_6$ :**  $H_2 \prec H_1 \prec P$ , similar to event  $e_4$ ,  $H_2$  and  $H_1$  fail as spares before  $P$  fails, similar to one spare SSP gate, where it is guaranteed that  $P$  does not fail during  $(0, \tau_2]$ . The

probability that the SSP gate fails during  $(0, t]$  can be calculated as in Eq. (11).

$$\Pr_{H_1 \sim H_2 \sim P} (T \leq t) = \int_0^t \int_0^t (\lambda_p e^{-\lambda_p \tau_1}) (\lambda_{H_1} e^{-\lambda_{H_1} \tau_2}) (\lambda_{H_2} e^{-\lambda_{H_2} \tau_3}) d\tau_1 d\tau_2 d\tau_3 \quad (11)$$

Based on Eq. (2), the unreliability function of the SSP gate with two HSSs is given in Eq. (12) given that the reliability function is  $R(t) = 1 - U(t)$ .

$$U(t) = \Pr_{P \sim H_1 \sim H_2} (T \leq t) + \Pr_{P \sim H_2 \sim H_1} (T \leq t) + \Pr_{H_2 \sim P \sim H_1} (T \leq t) + \Pr_{H_1 \sim H_2 \sim P} (T \leq t) + \Pr_{H_2 \sim H_1 \sim P} (T \leq t) + \Pr_{H_1 \sim P \sim H_2} (T \leq t) \quad (12)$$

### B. Reliability Function Verification Using CTMC

We use a CTMC model to formally verify the correctness of the reliability function  $R(t)$  derived in the previous section. Fig. 4 shows the CTMC model corresponding to the SSP gate with two HSSs illustrated in Fig. 1.

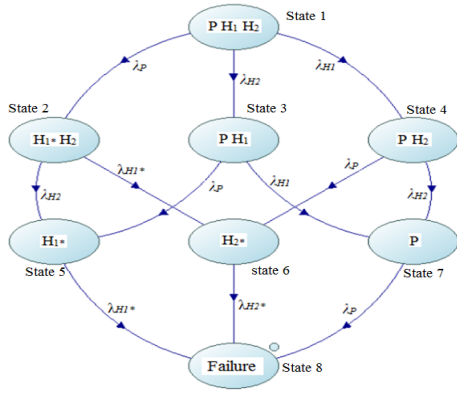


Fig. 4. The CTMC model of the SSP gate in Fig. 1

There are 8 states in the model, denoted as  $PH_1H_2$ ,  $H_1^*H_2$ ,  $PH_1$ ,  $PH_2$ ,  $H_1^*$ ,  $H_2^*$ ,  $P$ , and *Failure*. Each state holds the name of the surviving components, except the *Failure* state, which is the unavailability state. The reliability of the SSP gate is the sum of the probability of being in all available states, namely State 1 to State 7. Let  $P_i(t)$  be the probability of the system in state  $i$  at time  $t$ , where  $1 \leq i \leq 8$ , and  $P_{ij}(dt) = P[X(t+dt) = j | X(t) = i]$  be the incremental transition probability with random variable  $X(t)$ . The matrix  $[P_{ij}(dt)]$  defined in Eq. (13), where  $1 \leq i, j \leq 8$ , is the incremental one-step transition matrix of the CTMC defined in Fig. 4.

$$\begin{bmatrix} 1 - (\lambda_p + \lambda_{H_1} + \lambda_{H_2})dt & \lambda_p dt & \lambda_{H_2} dt & \lambda_{H_1} dt & 0 & 0 & 0 & 0 \\ 0 & 1 - (\lambda_{H_1} + \lambda_{H_2})dt & 0 & 0 & \lambda_{H_2} dt & \lambda_{H_1} dt & 0 & 0 \\ 0 & 0 & 1 - (\lambda_p + \lambda_{H_1})dt & 0 & \lambda_p dt & 0 & \lambda_{H_1} dt & 0 \\ 0 & 0 & 0 & 1 - (\lambda_p + \lambda_{H_2})dt & 0 & \lambda_p dt & \lambda_{H_2} dt & 0 \\ 0 & 0 & 0 & 0 & 1 - \lambda_{H_1} dt & 0 & 0 & \lambda_{H_1} dt \\ 0 & 0 & 0 & 0 & 0 & 1 - \lambda_{H_2} dt & 0 & \lambda_{H_2} dt \\ 0 & 0 & 0 & 0 & 0 & 0 & 1 - \lambda_p dt & \lambda_p dt \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 \end{bmatrix} \quad (13)$$

The transition matrix is a stochastic matrix with each row sums to 1, and it defines the probability for each state either remaining (when  $i = j$ ) or transiting to a different state (when  $i \neq j$ ) during the time interval  $dt$ . Given the initial probabilities of the states, the matrix can be used to describe the state transition process completely. From Eq. (13), we can derive the following relations as in Eqs. (14.1-14.7).

$$P_1(t+dt) = (1 - (\lambda_p + \lambda_{H_1} + \lambda_{H_2})dt) P_1(t) \quad (14.1)$$

$$P_2(t+dt) = \lambda_p dt P_1(t) + (1 - (\lambda_{H_1} + \lambda_{H_2})dt) P_2(t) \quad (14.2)$$

$$P_3(t+dt) = \lambda_{H_2} dt P_1(t) + (1 - (\lambda_p + \lambda_{H_1})dt) P_3(t) \quad (14.3)$$

$$P_4(t+dt) = \lambda_{H_1} dt P_1(t) + (1 - (\lambda_p + \lambda_{H_2})dt) P_4(t) \quad (14.4)$$

$$P_5(t+dt) = \lambda_{H_2} dt P_2(t) + \lambda_p dt P_3(t) + (1 - \lambda_{H_1} dt) P_5(t) \quad (14.5)$$

$$P_6(t+dt) = \lambda_{H_1} dt P_2(t) + \lambda_p dt P_4(t) + (1 - \lambda_{H_2} dt) P_6(t) \quad (14.6)$$

$$P_7(t+dt) = \lambda_{H_1} dt P_3(t) + \lambda_{H_2} dt P_4(t) + (1 - \lambda_p dt) P_7(t) \quad (14.7)$$

We derive a set of linear first-order differential equations as in Eqs. (15.1-15.7), which are state equations of the CTMC model assuming the initial probabilities  $P_1(0) = 1$ , and  $P_2(0) = P_3(0) = P_4(0) = P_5(0) = P_6(0) = P_7(0) = 0$ .

$$\frac{P_1(t+dt) - P_1(t)}{dt} = P_1'(t) = -(\lambda_p + \lambda_{H_1} + \lambda_{H_2})P_1(t) \quad (15.1)$$

$$\frac{P_2(t+dt) - P_2(t)}{dt} = P_2'(t) = \lambda_p P_1(t) - (\lambda_{H_1} + \lambda_{H_2})P_2(t) \quad (15.2)$$

$$\frac{P_3(t+dt) - P_3(t)}{dt} = P_3'(t) = \lambda_{H_2} P_1(t) - (\lambda_p + \lambda_{H_1})P_3(t) \quad (15.3)$$

$$\frac{P_4(t+dt) - P_4(t)}{dt} = P_4'(t) = \lambda_{H_1} P_1(t) - (\lambda_p + \lambda_{H_2})P_4(t) \quad (15.4)$$

$$\frac{P_5(t+dt) - P_5(t)}{dt} = P_5'(t) = \lambda_{H_2} P_2(t) + \lambda_p P_3(t) - \lambda_{H_1} P_5(t) \quad (15.5)$$

$$\frac{P_6(t+dt) - P_6(t)}{dt} = P_6'(t) = \lambda_{H_1} P_2(t) + \lambda_p P_4(t) - \lambda_{H_2} P_6(t) \quad (15.6)$$

$$\frac{P_7(t+dt) - P_7(t)}{dt} = P_7'(t) = \lambda_{H_1} P_3(t) + \lambda_{H_2} P_4(t) - \lambda_p P_7(t) \quad (15.7)$$

Using Laplace transformation to both sides of Eqs. (15.1-15.7) to derive Eqs. (16.1-16.7).

$$sP_1(s) - P_1(0) = -(\lambda_p + \lambda_{H_1} + \lambda_{H_2})P_1(s) \quad (16.1)$$

$$sP_2(s) - P_2(0) = \lambda_p P_1(s) - (\lambda_{H_1} + \lambda_{H_2})P_2(s) \quad (16.2)$$

$$sP_3(s) - P_3(0) = \lambda_{H_2} P_1(s) - (\lambda_p + \lambda_{H_1})P_3(s) \quad (16.3)$$

$$sP_4(s) - P_4(0) = \lambda_{H_1} P_1(s) - (\lambda_p + \lambda_{H_2})P_4(s) \quad (16.4)$$

$$sP_5(s) - P_5(0) = \lambda_{H_2} P_2(s) + \lambda_p P_3(s) - \lambda_{H_1} P_5(s) \quad (16.5)$$

$$sP_6(s) - P_6(0) = (\lambda_{H_1} P_2(s) + \lambda_p P_4(s) - \lambda_{H_2} P_6(s)) \quad (16.6)$$

$$sP_7(s) - P_7(0) = \lambda_{H_1} P_3(s) + \lambda_{H_2} P_4(s) - \lambda_p P_7(s) \quad (16.7)$$

Substituting the initial probabilities  $P_i(0)$ , where  $1 \leq i \leq 7$ , into Eqs. (16.1-16.7), we can derive the equations for  $P_1(s)$ ,  $P_2(s)$ ,  $P_3(s)$ ,  $P_4(s)$ ,  $P_5(s)$ ,  $P_6(s)$  and  $P_7(s)$ . By applying *inverse* Laplace transformation, we can solve the original linear first-order differential equations as follows.

$$P_1(s) = \frac{1}{(s + \lambda_p + \lambda_{H_1} + \lambda_{H_2})} \Rightarrow P_1(t) = e^{-(\lambda_p + \lambda_{H_1} + \lambda_{H_2})t}$$

$$P_2(s) = \frac{\lambda_p \cdot P_1(s)}{(s + \lambda_{H_1} + \lambda_{H_2})} \Rightarrow P_2(t) = \frac{\lambda_p}{\lambda_p + \lambda_{H_1} - \lambda_{H_2}} (e^{-(\lambda_{H_1} + \lambda_{H_2})t} - e^{-(\lambda_p + \lambda_{H_1} + \lambda_{H_2})t})$$

$$P_3(s) = \frac{\lambda_{H_2} \cdot P_1(s)}{(s + \lambda_p + \lambda_{H_1})} \Rightarrow P_3(t) = e^{-(\lambda_p + \lambda_{H_1})t} - e^{-(\lambda_p + \lambda_{H_1} + \lambda_{H_2})t}$$

$$P_4(s) = \frac{\lambda_{H_1} \cdot P_1(s)}{(s + \lambda_p + \lambda_{H_2})} \Rightarrow P_4(t) = e^{-(\lambda_p + \lambda_{H_2})t} - e^{-(\lambda_p + \lambda_{H_1} + \lambda_{H_2})t}$$

$$P_5(s) = \frac{\lambda_{H_1} \cdot P_2(s)}{(s + \lambda_{H_1}^*)} + \frac{\lambda_P \cdot P_3(s)}{(s + \lambda_{H_1}^*)}$$

$$\Rightarrow P_5(t) = \left[ \frac{\lambda_P}{\lambda_P + \lambda_{H_1} - \lambda_{H_1}^*} (e^{-\lambda_{H_1}^* t}) \right. \\ \left. - \frac{-\lambda_P}{\lambda_P + \lambda_{H_1} - \lambda_{H_1}^*} (e^{-\lambda_{H_1}^* t} + e^{-(\lambda_P + \lambda_{H_1}) t}) \right] \\ + \left[ \frac{\lambda_P (\lambda_P + \lambda_{H_1} - \lambda_{H_1}^* - \lambda_{H_1}^*)}{(\lambda_P + \lambda_{H_1} + \lambda_{H_2} - \lambda_{H_1}^*) (\lambda_P + \lambda_{H_1} - \lambda_{H_1}^*)} e^{-(\lambda_P + \lambda_{H_1} + \lambda_{H_2}) t} \right]$$

$$P_6(s) = \frac{\lambda_{H_1} \cdot P_2(s)}{(s + \lambda_{H_2}^*)} + \frac{\lambda_P \cdot P_4(s)}{(s + \lambda_{H_2}^*)}$$

$$\Rightarrow P_6(t) = \left[ \frac{(\lambda_P \cdot \lambda_{H_1}^*) (\lambda_P + \lambda_{H_2} - \lambda_{H_2}^*) + (\lambda_P \cdot \lambda_{H_1}) (\lambda_{H_2} + \lambda_{H_1}^* - \lambda_{H_2}^*)}{(\lambda_P + \lambda_{H_1} + \lambda_{H_2} - \lambda_{H_2}^*) (\lambda_{H_2} + \lambda_{H_1}^* - \lambda_{H_2}^*) (\lambda_P + \lambda_{H_2} - \lambda_{H_2}^*)} e^{-(\lambda_{H_2}^*) t} \right. \\ \left. + \left[ \frac{\lambda_P \cdot \lambda_{H_1}^*}{(\lambda_P + \lambda_{H_1} - \lambda_{H_1}^*) (\lambda_{H_2}^* - \lambda_{H_2} - \lambda_{H_1}^*)} (e^{-\lambda_{H_2}^* t} + e^{-(\lambda_P + \lambda_{H_1}) t}) \right] \right. \\ \left. + \left[ \frac{\lambda_P}{-\lambda_P - \lambda_{H_2} + \lambda_{H_2}^*} (e^{-\lambda_P + \lambda_{H_2}} t) \right] \right. \\ \left. + \left[ \frac{\lambda_P (\lambda_P + \lambda_{H_1})}{(\lambda_P + \lambda_{H_1} + \lambda_{H_2} - \lambda_{H_2}^*) (\lambda_P + \lambda_{H_1} - \lambda_{H_1}^*)} e^{-(\lambda_P + \lambda_{H_1} + \lambda_{H_2}) t} \right] \right]$$

$$P_7(s) = \frac{\lambda_{H_1} \cdot P_3(s)}{(s + \lambda_P)} + \frac{\lambda_{H_2} \cdot P_4(s)}{(s + \lambda_P)}$$

$$\Rightarrow P_7(t) = e^{-(\lambda_P) t} - e^{-(\lambda_P + \lambda_{H_1}) t} - e^{-(\lambda_P + \lambda_{H_2}) t} + e^{-(\lambda_P + \lambda_{H_1} + \lambda_{H_2}) t}$$

The reliability function  $R(t)$  from CTMC analysis is given as in Eq. (17).

$$R(t) = P_1(t) + P_2(t) + P_3(t) + P_4(t) + P_5(t) + P_6(t) + P_7(t) \quad (17)$$

We compute the system reliabilities using the reliability function  $R(t)$  from both the proposed approach and the CTMC approach presented in Eq. (12) and Eq. (17), respectively, in Table 1. The results show they are perfectly matched.

Table 1. R(t) analysis results - proposed method vs. CTMC

Time (days)	R(t) - proposed method	R(t) - CTMC
90	0.9815	0.9815
180	0.9019	0.9019
300	0.7299	0.7299
1000	0.031	0.031

## V. CASE STUDY

In this case study, we show how to model and analyze the reliability of a cloud-based system with two HSSs for each critical component using extended DFT, and then estimate rejuvenation schedules based on reliability quantitative analysis generated by our proposed approach in Section 4.

Figure 5 shows a cloud-based system that consists of an application server  $PA$  and a database server  $PB$ . To enhance the system reliability, four hot spare components  $HA_1$ ,  $HA_2$  are set up for  $PA$ , and  $HB_1$  and  $HB_2$  are set up for  $PB$ . The four HSSs are ready to take over the workload if the primary ones fail. The case study shows the reliability analysis applicable to SSP gate with two HSSs for each primary component. We set the reliability threshold to 0.99 as a minimum requirement for system reliability. For this case study, we assume constant failure rates for the servers, where  $\lambda_{PA} = 0.004/\text{day}$ ,  $\lambda_{HA_1} = \lambda_{HA_2}$

$= 0.0025/\text{day}$ ,  $\lambda_{PB} = 0.005/\text{day}$ ,  $\lambda_{HB_1} = \lambda_{HB_2} = 0.003/\text{day}$ , using the same failure rates as in previous work [6], so the obtained results can be readily compared.

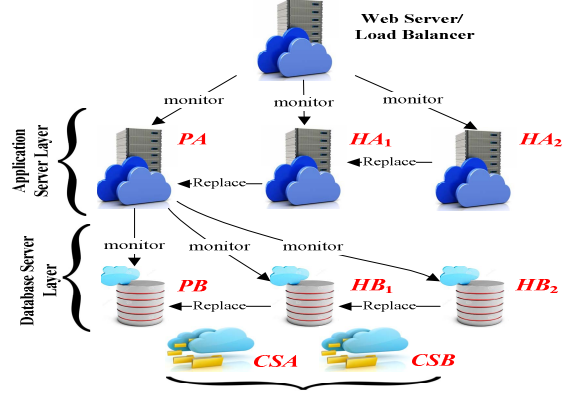


Fig. 5. A cloud-based system with software spares

As stated earlier, the failure rates of the HSS servers are lower than those of their corresponding primary ones because HSSs are not subject to the same workload; thus they have no software aging issues, and less likely to fail. Yet, when a primary server fails, the failure rate of a substituting HSS increases since it assumes the primary component workload, i.e.,  $\lambda_{PA} = \lambda_{HA_1}^* = \lambda_{HA_2}^* = 0.004$ , and  $\lambda_{PB} = \lambda_{HB_1}^* = \lambda_{HB_2}^* = 0.005$ .

The case study also involves CSS components, namely  $CSA$  and  $CSB$ , which are used in the rejuvenation process. Note that a CSS is a stored image of a deployed VM instance that can be easily duplicated, thus only one CSS is needed for each of the primary and HSS components. In addition, since a CSS is stored as an image, its failure rate is considered to be 0. However, once a CSS component is duplicated and deployed, it will assume the failure rate of its corresponding role, either as a running primary component or as an HSS. The DFT model of the cloud-based software system for Phase 1 is shown in Fig. 6.

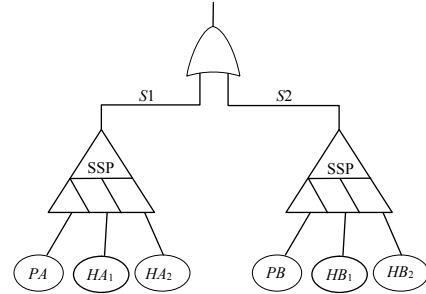


Fig. 6. DFT model of the cloud-based system - Phase 1

Since the system fails when either the application server or the database server fails, the two SSP gates are connected by an OR-gate. The reliability function of the OR-gate is derived using sum of disjoint product as in Eq. (18).

$$R(t) = 1 - U_{OR}(t) = 1 - (U_{S1}(t) + (1 - U_{S1}(t)) * U_{S2}(t)) \quad (18)$$

where  $U_{S1}(t)$  and  $U_{S2}(t)$  are the unreliability functions of the subtrees  $S1$  and  $S2$  that can be calculated using Eq. (12). We consider both scenarios mentioned in Section 3 for Phase 2 analysis. Fig. 7 represents the DFT model of the cloud-based system in Phase 2 for Scenario 1.



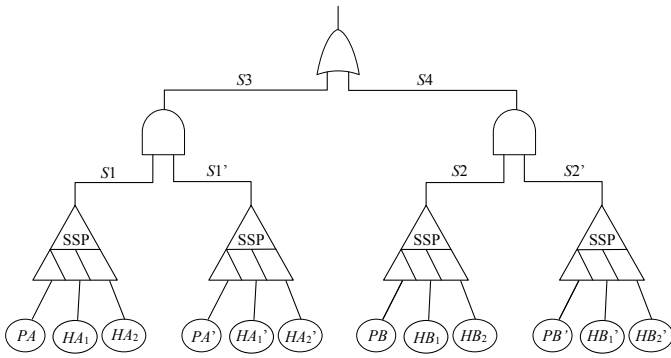


Fig. 7. DFT model of the cloud-based system - Phase 2 (Scenario1)

Similar to Phase 1, we can analyze the DFT model for Phase 2 Scenario 1 by splitting it into subtree sections. Starting from bottom to top, the unreliabilities  $U_{S1}(t)$ ,  $U_{S1'}(t)$ ,  $U_{S2}(t)$  and  $U_{S2'}(t)$  can be derived using Eq. (12).  $U_{S3}(t)$  and  $U_{S4}(t)$  can be calculated using the sum of disjoint product method for AND-gate shown in Eqs. (19-20). Finally, the system reliability for the OR-gate is derived as in Eq. (18) for Phase 1.

$$U_{S3}(t) = U_{S1}(t) * U_{S1'}(t) \quad (19)$$

$$U_{S4}(t) = U_{S2}(t) * U_{S2'}(t) \quad (20)$$

Moving to Phase 2 Scenario 2, the DFT model of the system is illustrated in Fig. 8 for the subsystem rejuvenation of the application servers.

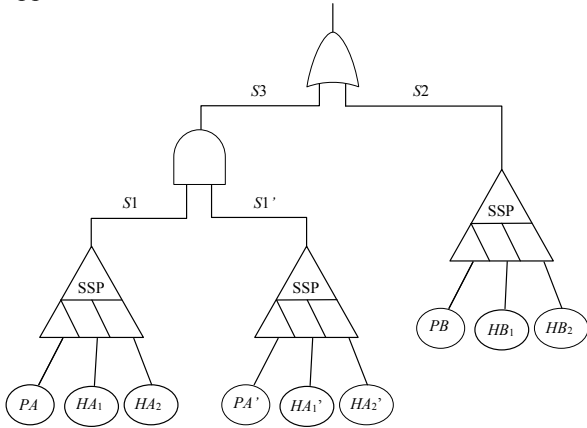


Fig. 8. DFT model of the cloud-based system - Phase 2 (Scenario2)

Using the same methodology for DFT analysis, we have the following subtrees  $U_{S1}(t)$ ,  $U_{S1'}(t)$ ,  $U_{S2}(t)$  and  $U_{S3}(t)$  in the DFT model. The unreliabilities  $U_{S1}(t)$ ,  $U_{S1'}(t)$  and  $U_{S2}(t)$  can be derived using Eq. (12).  $U_{S3}(t)$  is calculated using the sum of disjoint product method for AND-gate shown in Eq. (19).

As we have shown how to derive the system reliability in both Phase 1 and Phase 2, including the two different scenarios, the next step is to show the difference and the impact of employing 2-HSSs vs. 1-HSS [6] in terms of reliability and rejuvenation scheduling in a cloud-based system. In addition, we study the impacts of using Scenario 1 vs. Scenario 2 for rejuvenation scheduling for a cloud-based system with multiple HSSs subject to software aging.

Figure 9 illustrates the details of the difference between the two cases based on Scenario 1. Note that 1-HSS results are formerly provided in [6]. From the figure, we can see that the system reliability is kept very high during the transition. According to Fig. 9, the reliability threshold for 2-HSSs is reached at 48 days, hence it is suggested that the system should be rejuvenated every 48 days under Scenario 1. On the other hand, we can also see that the system needs to be rejuvenated every 18 days with a single HSS usage. Comparing rejuvenation scheduling based on reliability analysis for both cases over the period of 120 days, we notice that the system with 2-HSSs only needs two rejuvenations (at 48 and 96 days), but it requires six rejuvenations with a single HSS for its critical component. Therefore, Scenario 1 with 2-HSSs results in  $(6*2-6*2)/(6*2) = 66\%$  reduction in cost and management for software rejuvenation, while keeping the system above the same reliability threshold (0.99).

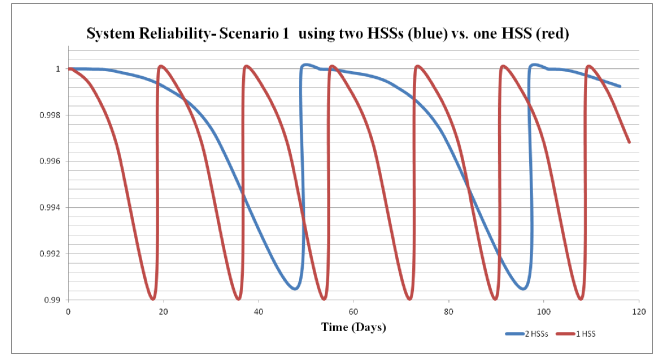


Fig. 9. Rejuvenation scheduling: 2-HSS vs. 1-HSS (Scenario1)

Figure 10 shows Scenario 2 for component-specific software rejuvenation. According to the figure, when the system reliability reaches the threshold in 48 days, the components with the lowest reliability, i.e., the database servers, are scheduled for rejuvenation first.

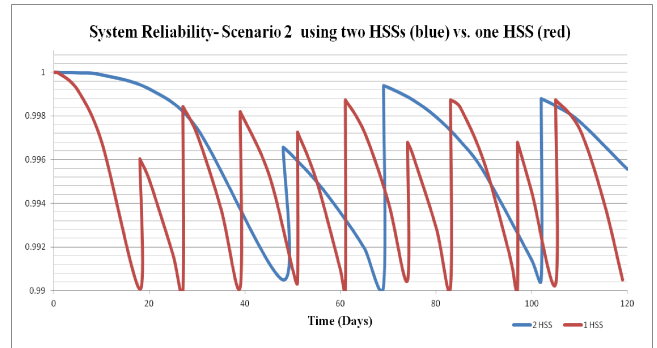


Fig. 10. Rejuvenation scheduling: 2-HSS vs. 1-HSS (Scenario2)

The rejuvenation induces a partial spike in the reliability curve, and then the system reliability is continuously monitored until it reaches the threshold again at the 69<sup>th</sup> day. At this point, the application server components become the ones with the lowest reliability. As a result, there will be an alternation in rejuvenation process for the two subsystems. We can see three rejuvenations for Scenario 2 with 2-HSSs vs. nine rejuvenations for 1-HSS design. Therefore, Scenario 2 with 2-HSSs results in  $(9-3)/(9) = 66\%$  reduction in cost and

management for software rejuvenation, while keeping the system above the same reliability threshold (0.99).

Figure 11 compares the two scenarios with two HSSs in 120 days. Scenario 1 has two rejuvenations that require us to rejuvenate both of the application and database servers. On the other hand, Scenario 2 has three rejuvenations that only require us to rejuvenate either the application servers or the database servers each time. Thus, by using Scenario 2, we can reduce the rejuvenation cost and management by  $(2*2-3)/4 = 25\%$  compared to the case of Scenario 1.

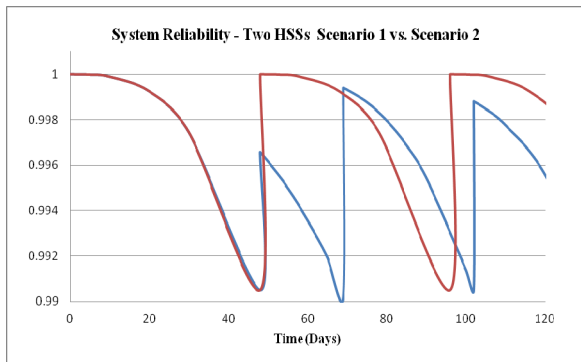


Fig. 11. Rejuvenation scheduling for 2-HSS: Scenario 1 vs. Scenario 2

## VI. CONCLUSIONS AND FUTURE WORK

In this paper, we introduced a reliability-based approach using two HSSs for critical components during normal running time in cloud-based software systems. We defined an extension of DFT, called SSP gate, which can be used to evaluate the reliability of a cloud-based system with multiple software spares for its critical components. Our approach has been verified using CTMC for constant failure rates. The case study showed that using the proposed approach, a rejuvenation schedule can be derived to maintain the system reliability of a trusted cloud-based software system with multiple software spare components above a certain level.

For future work, a measurement-based approach can be adopted for collecting empirical data relative to the software aging phenomenon, and then we can use data fitting technique to obtain the *pdfs* of the critical software components. Once the *pdfs* become available, they can be plugged into our proposed analytical approach to evaluate the system reliability and estimate the rejuvenation schedules based on the collected data. Finally, we envision modeling and analyzing cloud-based systems with active standby spare components, which can share workload with the primary ones, as a future, and more ambitious research direction.

## REFERENCES

- [1] K. V. Vishwanath and N. Nagappan, "Characterizing cloud computing hardware reliability," in *Proc. of the 1st ACM symposium on Cloud Computing (SoCC'10)*, Indianapolis, IN, USA, June 10-11, 2010, pp. 193-204.
- [2] M. Grotte, R. Matias, and K. S. Trivedi, "The fundamentals of software aging," in *Proc. of the 1st Int'l Workshop on Software Aging and Rejuvenation (WoSAR 2008)*, ISSRE, Seattle, WA, USA, November 11-14, 2008, pp. 1-6.
- [3] H. Pham, *System Software Reliability*, Springer Series in Reliability Engineering, Springer-Verlag, London, 2006.

- [4] M. Grotte, R. Matias, and K. S. Trivedi, "The fundamentals of software aging," in *Proc. of the 1st International Workshop on Software Aging and Rejuvenation (WoSAR 2008)*, ISSRE, Seattle, WA, USA, November 11-14, 2008, pp. 1-6.
- [5] Y. Huang, C. Kintala, N. Kolettis, and N. Fulton, "Software rejuvenation: analysis, module and applications," in *Proc. of the Twenty-Fifth International Symposium on Fault-Tolerant Computing (FTCS '95)*, Pasadena, CA, USA, June 27-30, 1995, pp. 381-390.
- [6] J. Rahme and H. Xu, "A software reliability model for cloud-based software rejuvenation using dynamic fault trees," *Int'l Journal of Software Engineering and Knowledge Engineering (IJSEKE)*, vol. 25, nos. 9 & 10, 2015, pp. 1491-1513.
- [7] V. Castelli, R. E. Harper, and P. Heidelberger, *et al.*, "Proactive management of software aging," *IBM Journal of Research and Development*, vol. 45, no. 2, 2001, pp. 311-332.
- [8] M. Grottke, L. Li, K. Vaidyanathan, and K. S. Trivedi, "Analysis of software aging in a web server," *IEEE Trans. on Reliability*, vol. 55, no. 3, 2006, pp. 411-420.
- [9] F. Machida, A. Andrzejak, R. Matias and E. Vicente, "On the effectiveness of Mann-Kendall test for detection of software aging," in *Proc. of the IEEE International Symposium on Software Reliability Engineering Workshops (ISSREW)*, Pasadena, CA, USA, November 4-7, 2013, pp. 269-274.
- [10] J. Guo, Y. Ju, Y. Wang, and X. Li, "The prediction of software aging trend based on user intention," in *Proc. of the IEEE Youth Conference on Information Computing and Telecommunications (YC-ICT)*, Beijing, China, November 28-30, 2010, pp. 206-209.
- [11] D. Cotroneo, R. Natella, and R. Pietrantuono, "Is software aging related to software metrics?" in *Proc. of the IEEE Second Int'l Workshop on Software Aging and Rejuvenation (WoSAR)*, San Jose, CA, USA, November 2, 2010, pp. 1-6.
- [12] A. Bobbio, M. Sereno, and C. Anglano, "Fine grained software degradation models for optimal rejuvenation policies," *Performance Evaluation*, vol. 46, no. 1, 2001, pp. 45-62.
- [13] K. Vaidyanathan, D. Selvamuthu, and K. S. Trivedi, "Analysis of inspection-based preventive maintenance in operational software systems," in *Proc. of the 21st IEEE Symp. on Reliable Distributed Systems (SRDS 2002)*, Suita, Japan, October 13-16, 2002, pp. 286-295.
- [14] V. P. Koutras and A. N. Platis, "Applying software rejuvenation in a two node cluster system for high availability," in *Proc. of the Int'l Conference on Dependability of Computer Systems*, Szklarska, Poreba, May 25-27, 2006, pp. 175-182.
- [15] J. Ke, Z. Su, K. Wang, and Y. Hsu, "Simulation inferences for an availability system with general repair distribution and imperfect fault coverage," *Simulation Modelling Practice and Theory*, vol. 18, no. 3, 2010, pp. 338-347.
- [16] T. Zhang, M. Xie, and M. Horigome, "Availability and reliability of k-out-of-(M+N):G warm standby systems," *Reliability Engineering and System Safety*, vol. 91, no. 4, 2006, pp. 381-387.
- [17] S. Distefano, F. Longo, and M. Scarpa, "Availability assessment of HA standby redundant clusters," in *Proc. 29th IEEE Int'l Symp. Reliable Distributed Systems*, 2010, pp. 265-274.
- [18] D. Liu, C. Zhang, W. Xing, R. Li, and H. Li, "Quantification of cut sequence set for fault tree analysis," in *HPCC2007, Lecture Notes in Computer Science*, 2007, Springer-Verlag, no. 4782, pp. 755-765.
- [19] A. B. Rauzy, "Sequence algebra, sequence decision diagrams and dynamic fault trees," *Reliability Engineering and System Safety*, vol. 96, no. 7, Jul. 2011, pp. 785-792.
- [20] G. Levitin, L. Xing, and Y. Dai, "Cold vs. hot standby mission operation cost minimization for 1-out-of-N systems," *European Journal of Operational Research*, vol. 234, no. 1, Apr. 2014, pp. 155-162.
- [21] F. Machida, D. Kim, and K. Trivedi, "Modeling and analysis of software rejuvenation in a server virtualized system," in *Proc. of the IEEE Second Int'l Workshop Software Aging and Rejuvenation (WoSAR)*, San Jose, CA, USA, Nov. 2, 2010, pp. 1-6.
- [22] T. Thein, S.-D. Chi, and J. S. Park, "Availability modeling and analysis on virtualized clustering with rejuvenation," *Int'l Journal of Computer Science and Network Security (IJCNNS)*, vol. 8, no. 9, 2008, pp. 72-80.